



Universal Outlier Hypothesis Testing

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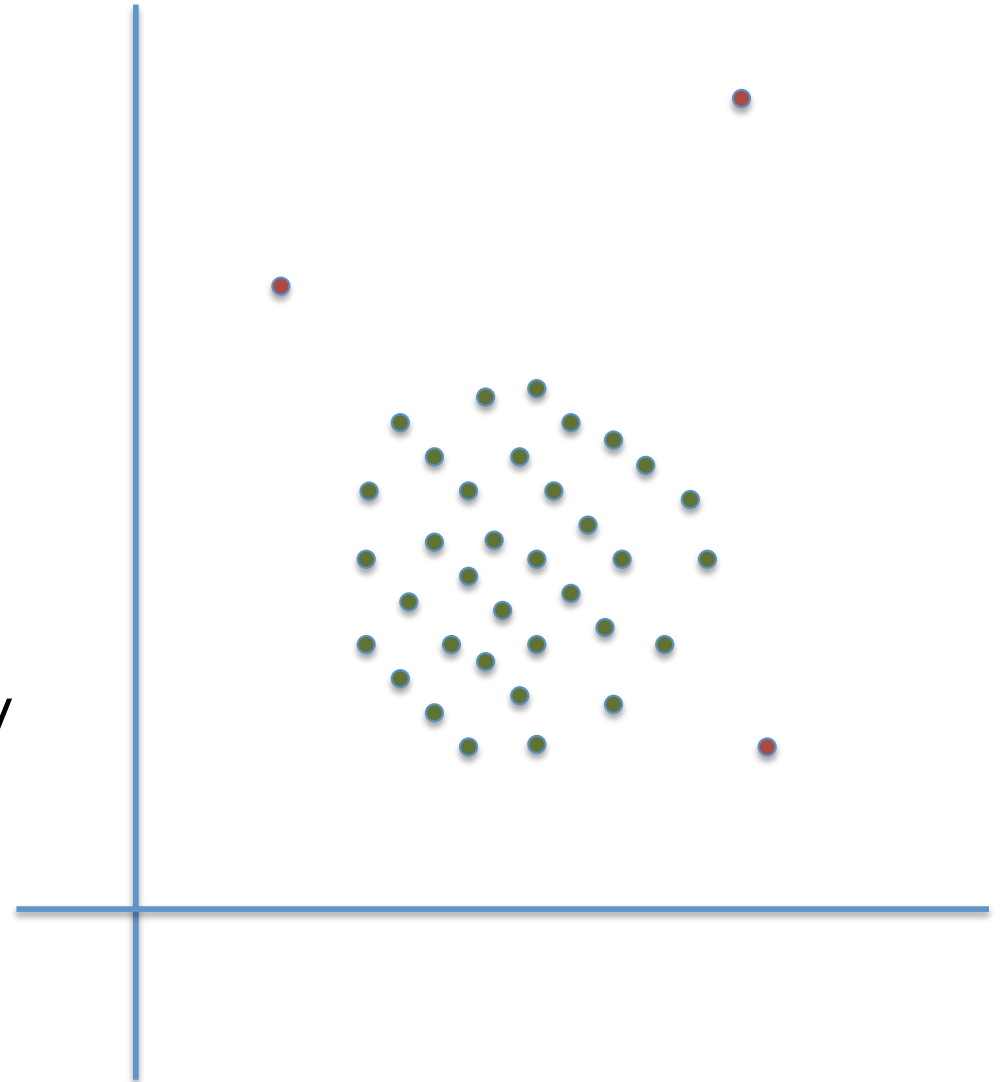
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(with Yun Li and Prof. Venu V. Veeravalli)



Statistical Outlier Detection

- Single sequence of observations
- Generic observations follow some fixed (possibly unknown) distribution or generating mechanism
- Outliers follow different generating mechanism
- Goal: To find outliers efficiently
- Applications: fraud detection, public health monitoring, cleaning up data





Fraud Detection



- **Example:** spending records for a male graduate student

Trans- actions	Grocery	Gas	---	Books	Grocery		
Amount	30\$	35\$	---	75\$	35\$		



Generic behavior



Fraud Detection



- **Example:** spending records for a male graduate student

Trans- actions	Grocery	Gas	---	Books	Grocery	Spa	Cosmetics
Amount	30\$	35\$	---	75\$	35\$	250\$	500\$

Generic behavior

Fraudulent behavior



Fraud Detection: Group Monitoring

- Male graduate students

Student 1	Student 2	Student 3	...	Student M
Grocery	Dining	Grocery	...	Gas
Dining	Grocery	Gas	...	Books
...
Books	Movie	Books	...	Grocery
Movie	Books	Grocery	...	Dining
Grocery	Gas	spa	...	Movie
Gas	Books	Cosmetics	...	Grocery



Outlier Hypothesis Testing

- M sequences of observations, with M large
- Almost all sequences are generated from common **typical** distribution
- Small subset of sequences generated from different (**outlier**) distribution



Outlier Hypothesis Testing

- M sequences of observations, with M large
- Almost all sequences are generated from common **typical** distribution
- Small subset of sequences generated from different (**outlier**) distribution
- **Special case:**
 - Exactly **one** sequence is generated from outlier distribution
 - **Goal:** to detect outlier sequence efficiently
 - **Universal setting:** neither typical nor outlier distributions known; no training data provided



Universal Outlier Hypothesis Testing

- Typical distribution π
- Outlier distribution μ

	1	2		M
H_1	μ	π	π	π
H_2	π	μ	π	π
	π	π	μ	π
	π	π	π	μ
H_M	π	π	π	μ



Applications: Outlier Hypothesis Testing

- Search problems and target tracking
- Sensor network applications: event detection, environment monitoring
- Fraud detection and anomaly detection in big data



Mathematical Model

$$\begin{array}{cccc} \boxed{\begin{array}{c} y_1^{(1)} \\ y_2^{(1)} \\ \vdots \\ y_n^{(1)} \end{array}} & \boxed{\begin{array}{c} y_1^{(2)} \\ y_2^{(2)} \\ \vdots \\ y_n^{(2)} \end{array}} & \dots & \boxed{\begin{array}{c} y_1^{(M)} \\ y_2^{(M)} \\ \vdots \\ y_n^{(M)} \end{array}} \\ y^{(1)} & y^{(2)} & & y^{(M)} \end{array}$$

$$H_i : p_i(y^{(1)}, \dots, y^{(M)}) = \prod_{k=1}^n \left[\mu(y_k^{(i)}) \prod_{j \neq i} \pi(y_k^{(j)}) \right]$$



Universal Outlier Hypothesis Testing

$$H_i : p_i(y^{Mn}) = \prod_{k=1}^n \left[\mu(y_k^{(i)}) \prod_{j \neq i} \pi(y_k^{(j)}) \right]$$

Nothing is known about (μ, π) except that they are distinct

Universal Test: $\delta : \mathcal{Y}^{Mn} \rightarrow \{1, \dots, M\}$



Universal Outlier Hypothesis Testing

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Nothing is known about (μ, π) except that they are distinct

Universal Test: $\delta : \mathcal{Y}^{Mn} \rightarrow \{1, \dots, M\}$

Independent of (μ, π)





Performance Metrics

- Maximal error probability:

$$e\left(\delta, \left(\mu, \pi\right)\right) = \max_i \mathbb{P}_i \left\{ \delta(\mathbf{y}^{Mn}) \neq i \right\}$$

- Exponent for maximal error probability:

$$\alpha\left(\delta, \left(\mu, \pi\right)\right) = \lim_{n \rightarrow \infty} -\frac{1}{n} \log e\left(\delta, \left(\mu, \pi\right)\right)$$



Performance Metrics

- Maximal error probability:

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- Exponent for maximal error probability:

$$\alpha\left(\delta, \left(\mu, \pi\right)\right) = \lim_{n \rightarrow \infty} -\frac{1}{n} \log e\left(\delta, \left(\mu, \pi\right)\right)$$

Consistency : $e \rightarrow 0$ as $n \rightarrow \infty$

Exponential Consistency : $\alpha > 0$



Background: Binary Hypothesis Testing

$$H_1 : p_1(\mathbf{y}) = \prod_{k=1}^n \pi(y_k) \quad H_2 : p_2(\mathbf{y}) = \prod_{k=1}^n \mu(y_k)$$

If (μ, π) known, $\delta_{\text{ML}}(\mathbf{y}) = \underset{i}{\operatorname{argmax}} \log p_i(\mathbf{y})$

has $\alpha(\delta, (\mu, \pi)) = C(\mu, \pi) > 0$



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$$\text{Chernoff Info : } C(\mu, \pi) = \max_{0 \leq s \leq 1} -\log \left(\sum_y \mu(y)^s \pi(y)^{1-s} \right)$$



Outlier Hypothesis Testing: Known (μ, π)

$$H_i : p_i(\mathbf{y}^{Mn}) = \prod_{k=1}^n \left[\mu(y_k^{(i)}) \prod_{j \neq i} \pi(y_k^{(j)}) \right]$$

ML Rule : $\delta_{\text{ML}}(\mathbf{y}^{Mn}) = \underset{i}{\operatorname{argmax}} \log p_i(\mathbf{y}^{Mn})$

	1	2			M
H_1	μ	π	π	π	π
H_2	π	μ	π	π	π
	π	π	μ	π	π
	π	π	π	μ	π
H_M	π	π	π	π	μ

Exponential Consistency : $\alpha(\delta_{\text{ML}}, (\mu, \pi)) = 2B(\mu, \pi)$

Bhattacharya Distance : $B(\mu, \pi) = -\log \left(\sum_y \mu(y)^{1/2} \pi(y)^{1/2} \right)$



Binary Hypothesis Testing: Unknown μ

$$H_1 : p_1(\mathbf{y}) = \prod_{k=1}^n \pi(y_k) \quad H_2 : p_2(\mathbf{y}) = \prod_{k=1}^n \mu(y_k)$$

If μ unknown

for any given δ there exists μ s.t. $\alpha = 0$

No exponential consistency!

Outlier Hypothesis Testing: Unknown μ

μ unknown : $\hat{\mu}_i = \gamma_i \leftarrow$ empirical distribution

$$H_i : \hat{p}_i(y^{Mn}) = \prod_{k=1}^n \left[\hat{\mu}_i(y_k^{(i)}) \prod_{j \neq i} \pi(y_k^{(j)}) \right]$$

Generalized Likelihood (GL) Rule :

$$\delta_{\text{GL}}(y^{Mn}) = \underset{i}{\operatorname{argmax}} \log \hat{p}_i(y^{Mn})$$

	1	2			M
H_1	μ	π	π	π	π
H_2	π	μ	π	π	π
	π	π	μ	π	π
	π	π	π	μ	π
H_M	π	π	π	π	μ

Exponential Consistency : $\alpha(\delta_{\text{GL}}, (\mu, \pi)) = 2B(\mu, \pi)$

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$$\delta_{\text{GL}}(y^{Mn}) = \underset{i}{\operatorname{argmax}} \log \hat{p}_i(y^{Mn})$$

	1	2	M		
H_1	μ	π	π	π	π
H_2	π	μ	π	π	π
	π	π	μ	π	π
	π	π	π	μ	π
H_M	π	π	π	π	μ

Exponential Consistency : $\alpha(\delta_{\text{GL}}, (\mu, \pi)) = 2B(\mu, \pi)$

Same as known μ, π



Sanov's Theorem

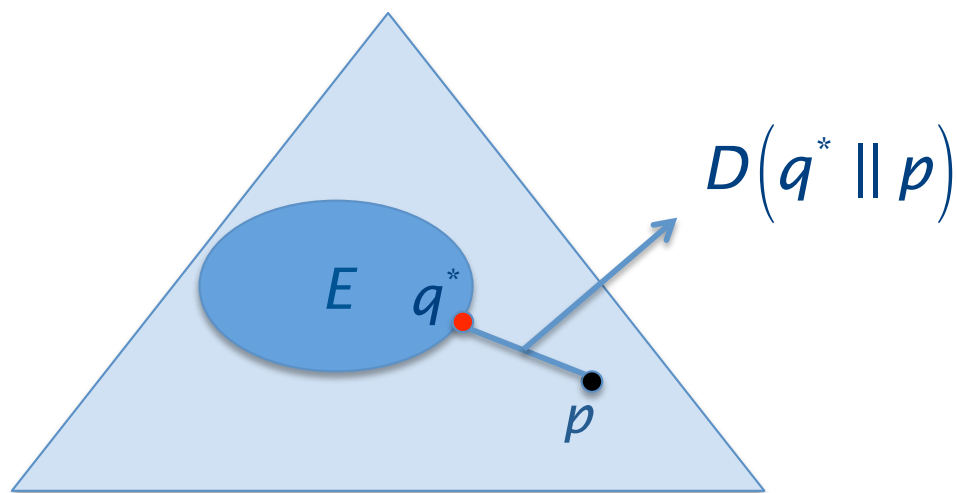
- **Sanov's Theorem:** For i.i.d. rvs $Y^n \sim p$, exponent of probability that random empirical distribution falls in closed set E is



Key Tool: Sanov's Theorem

- **Sanov's Theorem:** For i.i.d. rvs $Y^n \sim p$, exponent of probability that random empirical distribution falls in closed set E is

$$\lim_n -\frac{1}{n} \log \mathbb{P} \left\{ \text{Empirical}(Y^n) \in E \right\} = \min_{q \in E} D(q \parallel p)$$





Proposed Universal Test

(μ, π) not known : $\hat{\mu}_i = \gamma_i$ $\hat{\pi}_i = \frac{1}{M-1} \sum_{j \neq i} \gamma_j$

↑ ↑
empirical distributions

$$H_i : \quad \hat{\hat{p}}_i \left(\mathbf{y}^{Mn} \right) = \prod_{k=1}^n \left[\hat{\mu}_i \left(\mathbf{y}_k^{(i)} \right) \prod_{j \neq i} \hat{\pi}_i \left(\mathbf{y}_k^{(j)} \right) \right]$$

Generalized Likelihood (GL) Rule :

$$\delta_{\text{GL}}(\mathbf{y}^{Mn}) = \underset{i}{\operatorname{argmax}} \log \hat{\hat{p}}_i(\mathbf{y}^{Mn})$$

	1	2	...	M
H_1	μ	π	π	π
H_2	π	μ	π	π
\vdots	π	π	\ddots	π
H_M	π	π	π	μ



Proposed Universal Test

$$(\mu, \pi) \text{ not known : } \hat{\mu}_i = \gamma_i \quad \hat{\pi}_i = \frac{1}{M-1} \sum_{j \neq i} \gamma_j$$

$$H_i : \hat{p}_i(y^{Mn}) = \prod_{k=1}^n \left[\hat{\mu}_i(y_k^{(i)}) \prod_{j \neq i} \hat{\pi}_i(y_k^{(j)}) \right]$$

Generalized Likelihood (GL) Rule :

$$\delta_{GL}(y^{Mn}) = \underset{i}{\operatorname{argmax}} \log \hat{p}_i(y^{Mn})$$

$$= \underset{i}{\operatorname{argmin}} \sum_{j \neq i} D \left(\gamma_j \parallel \frac{1}{M-1} \sum_{k \neq i} \gamma_k \right)$$

	1	2		M
H_1	μ	π	π	π
H_2	π	μ	π	π
	π	π	μ	π
	π	π	π	μ
H_M	π	π	π	μ



Proposed Universal Test

$$(\mu, \pi) \text{ not known : } \hat{\mu}_i = \gamma_i \quad \hat{\pi}_i = \frac{1}{M-1} \sum_{j \neq i} \gamma_j$$

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Generalized Likelihood (GL) Rule :

$$\delta_{GL}(y^{Mn}) = \underset{i}{\operatorname{argmax}} \log \hat{p}_i(y^{Mn})$$

$$= \underset{i}{\operatorname{argmin}} \sum_{j \neq i} D \left(\gamma_j \parallel \frac{1}{M-1} \sum_{k \neq i} \gamma_k \right) \leftarrow \text{key statistic}$$

	1	2	M		
H ₁	μ	π	π	π	π
H ₂	π	μ	π	π	π
	π	π	μ	π	π
	π	π	π	μ	π
H _M	π	π	π	π	μ



Performance of Universal Test

$$\alpha\left(\delta, \left(\mu, \pi\right)\right) = \min_{q_1, \dots, q_M} D(q_1 \parallel \mu) + D(q_2 \parallel \pi) + \dots + D(q_M \parallel \pi)$$

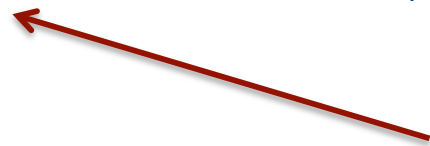
$$\sum_{j \neq 1} D(q_j \parallel \frac{1}{M-1} \sum_{k \neq 1} q_k) \geq \sum_{j \neq 2} D(q_j \parallel \frac{1}{M-1} \sum_{k \neq 2} q_k)$$



Performance of Universal Test

Universally exponential consistency!

$$\alpha\left(\delta, \left(\mu, \pi\right)\right) = \min_{q_1, \dots, q_M} D(q_1 \parallel \mu) + D(q_2 \parallel \pi) + \dots + D(q_M \parallel \pi)$$


$$> 0, \quad \forall (\mu, \pi)$$

$$\sum_{j \neq 1} D(q_j \parallel \frac{1}{M-1} \sum_{k \neq 1} q_k) \geq \sum_{j \neq 2} D(q_j \parallel \frac{1}{M-1} \sum_{k \neq 2} q_k)$$



Asymptotic Efficiency

- **Motivation:** When only π is known, optimal error exponent is $2B(\mu, \pi)$
- Estimate of π satisfies

$$\lim_{n \rightarrow \infty} \frac{1}{M} \sum_{i=1}^M \gamma_i = \frac{1}{M} \mu + \frac{M-1}{M} \pi,$$

$$\lim_{M \rightarrow \infty} \frac{1}{M} \mu + \frac{M-1}{M} \pi = \pi$$



Asymptotic Efficiency

Our universal outlier detector achieves error exponent lower bounded by

$$\min_{q: D(q||\pi) \leq \frac{1}{M-1}(2B(\mu, \pi) + C_\pi)} 2B(\mu, q)$$



Asymptotic Efficiency

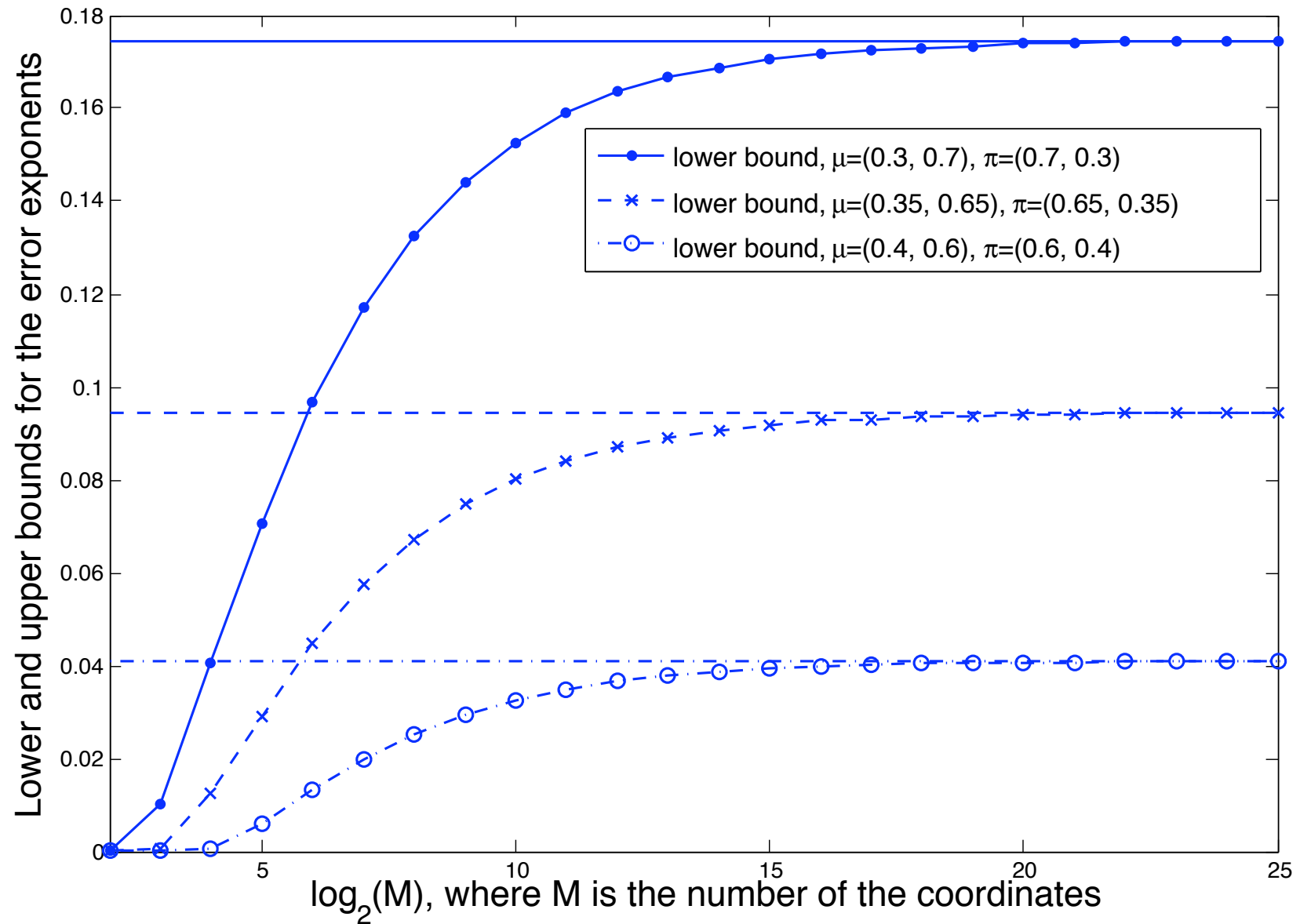
Our universal outlier detector achieves error exponent lower bounded by

$$\min_{q: D(q||\pi) \leq \frac{1}{M-1}(2B(\mu, \pi) + C_\pi)} 2B(\mu, q)$$

This lower bound is non-decreasing in $M \geq 3$,
and converges to $2B(\mu, \pi)$ as $M \rightarrow \infty$



Numerical Results





Extension to Multiple Outliers: Known (μ, π)

For $S \subseteq \{1, \dots, M\}$, $|S| = T$, T fixed and known

$$H_S : p_S(\mathbf{y}^{Mn}) = \prod_{k=1}^n \left[\prod_{i \in S} \mu(y_k^{(i)}) \prod_{j \notin S} \pi(y_k^{(j)}) \right]$$

ML Rule : $\delta_{\text{ML}}(\mathbf{y}^{Mn}) = \underset{S}{\operatorname{argmax}} \log p_S(\mathbf{y}^{Mn})$

Exponential Consistency : $\alpha(\delta_{\text{ML}}, (\mu, \pi)) = 2B(\mu, \pi)$



Proposed Universal Test: One Outlier

Generalized Likelihood (GL) Rule :

$$\begin{aligned}\delta_{\text{GL}}(y^{Mn}) &= \operatorname{argmax}_i \log \hat{p}_i(y^{Mn}) \\ &= \operatorname{argmin}_i \sum_{j \neq i} D\left(\gamma_j \parallel \frac{1}{M-1} \sum_{k \neq i} \gamma_k\right) \leftarrow \text{key statistic}\end{aligned}$$

	1	2			M
H_1	μ	π	π	π	π
H_2	π	μ	π	π	π
	π	π	μ	π	π
	π	π	π	μ	π
H_M	π	π	π	π	μ



Proposed Universal Test: One Outlier

Generalized Likelihood (GL) Rule :

$$\begin{aligned}\delta_{\text{GL}}(y^{Mn}) &= \operatorname{argmax}_i \log \hat{p}_i(y^{Mn}) \\ &= \operatorname{argmin}_i \sum_{j \neq i} D\left(\gamma_j \parallel \frac{1}{M-1} \sum_{k \neq i} \gamma_k\right) \leftarrow \text{key statistic}\end{aligned}$$

	1	2			M
H_1	μ	π	π	π	π
H_2	π	μ	π	π	π
	π	π	μ	π	π
	π	π	π	μ	π
H_M	π	π	π	π	μ

Summing over all typical
sequence indices under
hypothesis i

$\hat{\pi}_i$



Proposed Universal Test: Multiple Outliers

Generalized Likelihood (GL) Rule :

$$\delta_{\text{GL}}(y^{Mn}) = \underset{S}{\operatorname{argmin}} \sum_{j \notin S} D \left(\gamma_j \parallel \frac{1}{M-T} \sum_{k \notin S} \gamma_k \right)$$

Summing over all typical sequence indices under hypothesis S

$\hat{\pi}_S$



Asymptotic Efficiency

Our universal outlier detector achieves error exponent lower bounded by

$$q: D(q||\pi) \leq \frac{1}{M-T} (2B(\mu, \pi) + C_\pi) \quad \min_{2B(\mu, q)}$$



Asymptotic Efficiency

Our universal outlier detector achieves error exponent lower bounded by

$$\min_{q: D(q||\pi) \leq \frac{1}{M-T}(2B(\mu, \pi) + C_\pi)} 2B(\mu, q)$$

This lower bound is non-decreasing in M ,
and converges to $2B(\mu, \pi)$ as $M \rightarrow \infty$



Conclusion

- Generalized likelihood (GL) test is universally exponentially consistent for outlier hypothesis testing wherein number of outliers is fixed and known *a priori*
- GL test is asymptotically efficient in error exponent for large M even with no training data
- If number of outliers is not known *a priori*, there is no universally exponentially consistent test



Fraud Detection: Group Monitoring

- Male graduate students

Student 1	Student 2	Student 3	...	Student M
Grocery	Dining	Grocery	...	Gas
Dining	Grocery	Gas	...	Books
...
Books	Movie	Books	...	Grocery
Movie	Books	Grocery	...	Dining
Grocery	Gas	spa	...	Movie
Gas	Books	Cosmetics	...	Grocery